

Optimal Constrained Design of Control Charts Using Stochastic Approximations

Daniele Zago¹ Giovanna Capizzi¹ Peihua Qiu² 2023 INFORMS Annual Meeting October 15, 2023

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Introduction on Statistical Process Monitoring

• The goal of SPM is to detect the presence of a change in a sequential process

$$X_n \sim \begin{cases} F_0(\cdot) & \text{for } n = 0, 1, 2, \dots, \tau - 1 & \text{in-control (IC),} \\ F_1(\cdot) & \text{for } n = \tau, \tau + 1, 2, \dots & \text{out-of-control (OC).} \end{cases}$$

• Control charts are the main tools to test the stability of the process using incoming observations.

- 1. Monitoring statistic: C_n , e.g. $C_n = \max\left\{0, C_{n-1} + \left(\frac{X_n \mu_0}{\sigma_0}\right) k\right\}$.
- 2. Control limit: h > 0 for all n > 0.
- 3. **Run length**: $RL = \inf \{n : C_n > h\}$.

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Design of a control chart

- Selection of control chart tuning parameters, which typically depend on the expected process shift.
- Selection of the control limit h, which is usually chosen so that

 $ARL_{IC} = \mathbb{E}[RL|\tau = \infty] = ARL_0,$

for some value of ARL_0 (e.g. 200, 370, 500, ...)

Performance metrics

• Small values of $ARL_{\tau_0} = \mathbb{E}[RL|\tau = \tau_0]$ mean that the chart performs better.

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- Consider a CUSUM control chart $C_n = \max \left\{ 0, C_{n-1} + \left(\frac{X_n \mu_0}{\sigma_0} \right) k \right\}$ for monitoring changes in the mean of the process.
- Assume that the IC process observations are i.i.d. $X_n \sim \mathcal{N}(\mu_0, \sigma_0)$.
- Suppose that the expected OC mean is μ_1 and the (standardized) change to be detected is $\delta = \frac{\mu_1 \mu_0}{\sigma_0}$.
- Then, it is well-known that the tuning parameter that minimizes the ARL₁ is $k = \delta/2$, irrespective of the value of ARL₀.

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Optimization Using Stochastic Approximations

Chart design

- Control charts (and their run lengths) depend on a set of **tuning parameters** $\zeta \in \mathbb{R}^d$.
- Different values of tuning parameters allow better detection of different magnitudes of parameter shifts.
- The goal is to optimize the ARL_{OC} under the constraint on the in-control ARL,

$$\begin{split} \zeta^* &= \operatorname*{argmin}_{\zeta \in \mathcal{Z}} \mathbb{E}_1[\operatorname{RL}(\zeta, h(\zeta))] \\ \text{s.t. } \mathbb{E}_0[\operatorname{RL}(\zeta, h(\zeta))] &= \operatorname{ARL}_0, \end{split}$$

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Classical methods for optimization

Analytical methods

• Applicable only in very simple cases (CUSUM with Gaussian observations, ...)

Numerical methods

- Numerical quadrature (for example, Knoth, 2017).
- Limitations: applicable for some specific control charts, scales poorly when d > 1.

- Estimate ARL_{OC} for given ζ with a large number of simulations and use optimization tools.
 - Grid search (Qiu, 2008; Qiu and Xie, 2021).
 - Other numerical solvers (Capizzi and Masarotto, 2003; Mahmoud and Zahran, 2010).
- Limitations: function is treated as deterministic, methodologies are expensive to scale for d > 1 (as we will see later).

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Starting point

- Let $Q(\zeta, h(\zeta))$ be the **noisy** function we want to minimize
- Let $\Psi : \mathbb{R}^d \to \mathcal{Z}$ be the projection onto the nearest point in \mathcal{Z} .
- We would like to find the minimum of Q using a gradient descent iteration (Spall, 2003)

$$\widehat{\zeta}_{k+1} = \Psi\left(\widehat{\zeta}_k - a_k \frac{\partial Q(\zeta, h(\zeta))}{\partial \zeta}\Big|_{\zeta = \widehat{\zeta}}\right), \quad k = 1, 2, \dots$$
(1)

• Problem: The gradient in (1) is unknown and cannot be expressed in closed form.

- Estimate the gradient using finite differences and use $\widehat{\zeta}_{k+1} = \Psi\left(\widehat{\zeta}_k a_k \widehat{g}_k(\widehat{\zeta}_k)\right)$,
- When $\zeta \in \mathbb{R}^d$, the SA method requires evaluating the function Q at 2d parameter values.
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- Let $\Delta_k = (\Delta_{1k}, \dots, \Delta_{dk})$ be independent zero-mean random variables that are symmetric.
- Typically (Spall, 2003)

$$\Delta_{jk} \stackrel{\mathrm{iid}}{\sim} egin{cases} 1 & \mathrm{with \ probability \ } 1/2, \ -1 & \mathrm{with \ probability \ } 1/2. \end{cases}$$

- Perturb the current parameter estimates: $\widehat{\zeta}_k^+ = \Psi(\widehat{\zeta}_k + c_k \Delta_k)$ and $\widehat{\zeta}_k^- = \Psi(\widehat{\zeta}_k c_k \Delta_k)$.
- · Gradient estimate is

$$\widehat{\boldsymbol{g}}_{k}(\widehat{\boldsymbol{\zeta}}_{k}) = \frac{Q(\widehat{\boldsymbol{\zeta}}_{k}^{+}, h(\widehat{\boldsymbol{\zeta}}_{k}^{+})) - Q(\widehat{\boldsymbol{\zeta}}_{k}^{-}, h(\widehat{\boldsymbol{\zeta}}_{k}^{-}))}{2c_{k}} \begin{pmatrix} \Delta_{1k}^{-1} \\ \vdots \\ \Delta_{dk}^{-1} \end{pmatrix}, \qquad (2)$$

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Simultaneous perturbations approach

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Further enhancements

Polyak averaging

• The estimate of the optimum ζ^* at iteration k is the average over the optimization path,

$$\overline{\zeta}_{k} = \frac{1}{k - N_{\rm f}} \sum_{\ell = N_{\rm f} + 1}^{k} \widehat{\zeta}_{\ell}$$

• Averaging increases stability while providing a similar convergence rate to the solution.

Noise reduction

• Once $h(\widehat{\zeta}_k)$ is found, r = 100 simulations of $Q(\widehat{\zeta}_k)$ are used to estimate $\widehat{g}(\widehat{\zeta}_k)$.

Computational bottleneck

- Calculation of the control limits $h(\widehat{\zeta}_k^+)$ and $h(\widehat{\zeta}_k^-)$ is the algorithm's bottleneck.
- We use a low-precision SA algorithm (Capizzi and Masarotto, 2016) with a warm-start initialization that allows it to become more accurate as *k* increases.
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- We use a low-precision SA algorithm (Capizzi and Masarotto, 2016) with a warm-start initialization that allows it to become more accurate as *k* increases.

- A reasonable stopping rule of the algorithm is $|\mathbb{E}[\widehat{g}_{jk}(\widehat{\zeta}_k)]| \leq \nu$, for a small value of ν
- We leverage the asymptotic distribution of the PR averaging scheme (Maryak, 1997),

$$k^{1/3} \left[\mathcal{Q}'(\bar{\boldsymbol{\zeta}}_k) - \mathcal{Q}'(\boldsymbol{\zeta}^* - \boldsymbol{\mu}) \right] \stackrel{\sim}{\sim} N_d(\boldsymbol{0}, \mathcal{Q}''(\boldsymbol{\zeta}^* - \boldsymbol{\mu}) \boldsymbol{\Sigma} \mathcal{Q}''(\boldsymbol{\zeta}^* - \boldsymbol{\mu})^\top)$$

• Using a similar approach to Capizzi and Masarotto (2016), a stopping criterion can be defined as

$$\overline{N}_{s} = \inf \left\{ k > N_{\mathsf{m}} + N_{\mathsf{f}} : k \ge \left(\frac{z}{\nu}\right)^{2} \max_{j=1,\dots,p} \frac{1}{N - N_{\mathsf{f}}} \sum_{\ell=N_{\mathsf{f}}+1}^{k} \overline{g}_{j\ell}(\widehat{\zeta}_{k})^{2} \right\},$$

- *z* is the $[(1 \nu)/2]$ -th quantile of the standard normal distribution.
- N_m + N_f is specified to avoid a premature ending of the algorithm.

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"Deterministic" convergence criterion

• The "stochastic" convergence criterion \overline{N}_s is coupled with a classical "deterministic" convergence criterion,

$$\overline{N}_{a} = \inf \left\{ k > N_{\mathsf{m}} + N_{\mathsf{f}} : \|\overline{\zeta}_{k} - \overline{\zeta}_{k-1}\| < \varepsilon \right\}.$$

- The convergence criterion used in our simulation is $\overline{N} = \min \{\overline{N}_s, \overline{N}_a\}$.
- The convergence criteria \overline{N}_s and \overline{N}_a are "complementary":
 - \overline{N}_s is useful for optimizing "flat" functions, where the variance is small.
 - \overline{N}_a can help when functions have high curvature, because $\widehat{\zeta}_k$ will "jump around" the optimum and the effect will be averaged out in $\overline{\zeta}_k$ (Maryak, 1997).

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Numerical results

- At iteration k, divide each coordinate of a grid $[\zeta_{\min}^{(k)}, \zeta_{\max}^{(k)}]$ in m segments and calculate the objective function at the endpoints.
- Find the endpoint $\zeta^{(k+1)}$ with minimum value of the objective.
- The endpoints adjacent to $\zeta^{(k+1)}$ define $[\zeta^{(k+1)}_{min}, \zeta^{(k+1)}_{max}]$.

- Method based on reflection, extension, contraction, and shrinkage of a simplex.
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Results

- CUSUM control chart: $C_t = \max \{0, C_{t-1} + X_t k\}, X_t \sim N(\delta, 1)$ for various δ 's.
- Optimal parameter to detect the mean shift δ is $k^* = \delta/2$.



Figure 1: Estimated optimal parameter values over 100 optimizations.

Results



Figure 2: Computing times over 100 optimization.

Multidimensional scalability

• MEWMA control chart: $C_t = (I - \Lambda)C_{t-1} + \Lambda X_t$, with $\Lambda = \text{diag}(\lambda_1, \dots, \lambda_d)$ and $\lambda_j \in (0, 1)$ for all *j*. $X_{t,1} \sim \mathcal{N}(\delta, 1), X_{t,2} \sim \chi^2_{1+\delta\sqrt{2}}, X_{t,3} \sim \text{Pois}(1 + \delta)$.



Figure 3: Median, 0.1th and 0.9th quantiles of the computing times over 100 optimizations.

Extension to other metrics

Generalizations

- Sometimes, other metrics such as the median run length or quantiles of the run length are of interest (Knoth, 2015).
- The proposed SPSA algorithm is flexible enough to be generalized to other performance metrics based on the RL.

- Minimization of the out-of-control median run length with a constraint on the in-control median run length.
- *Q* becomes the median of the *r* simulated out-of-control run lengths.
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Conclusion

- We have proposed a novel methodology for a more **efficient design** of control charts tuning parameters.
- The only requirement is to be able to **simulate run lengths** from the IC process (e.g. using parametric/nonparametric bootstrap, bootstrap for time series, ...)
- The methodology is based on a **stochastic approximations** algorithm that is specifically designed for the constrained optimization problem.
- Numerical simulation suggest that the approach is **more efficient** than traditional approaches, especially for **multi-dimensional** tuning parameters.
- Due to its flexibility, it can be **generalized** to the optimization of various performance metrics such as median run length, run length quantiles, etc.

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Thank you for the attention

Questions
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Selection of the tuning constants in the SPSA algorithm i

- The proposed SPSA algorithm requires the selection of many tuning constants.
- A semi-automated way of selecting most of the tuning constants is available following the guidelines by Spall (1992), Spall (1998) and Spall (2003).
 - The gain sequences in the algorithm are defined as $a_k = a/(k + A + 1)^{\alpha}$ and $c_k = c/(k + 1)^{\beta}$, where α and β are pre-specified to be 0.602 and 0.101, respectively (Spall, 2003)
 - These gain sequences would result in a slow gain decay and ensure the convergence of $\widehat{\zeta}_k$ to ζ^* as $k \to \infty$ under some quite general assumptions, as proved by Spall (1992).
 - The constant *a*, *A* and *c* require a small preliminary adaptive step in order to be estimated.
 - *c* can be approximately set to be the standard deviation $\sigma_{\hat{\zeta}_0}$ of the OC RL calculated at the initial value $\hat{\zeta}_0$ (Spall, 1998).
 - In numerical studies, we have seen that setting c = min{σ_{ζ0}, 0.1} can avoid excessive perturbation of the tuning parameters in the early iterations.
 - *A* can be set to be 0.1 times the expected number of function evaluations. For example, the expected number of evaluations used in this paper is 150, resulting in $A = 0.1 \times 150 = 15$.

• (Spall, 1998) recommends selecting *a* to be the expected magnitude change in $\hat{\zeta}_k$ during the first few iterations. Specifically,

$$a=s\cdot(A+1)^{\alpha}/\overline{G},$$

where *s* is the initial step size and $\overline{G} = \frac{1}{d} \sum_{j=1}^{d} \sum_{l=1}^{n_c} \widehat{g}_{jl}(\widehat{\zeta}_0)/n_c$ is a preliminary estimate of the average value of the gradient in $\widehat{\zeta}_0$ based on n_c simulated RLs.

• For instance, a reasonable initial step size *s* for an EWMA chart could be 0.2, and setting $n_c = 20$ is found to be appropriate to estimate the gradient at the beginning of the algorithm.